

The Aitken Model

The Aitken Model (AM):

Suppose

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon},$$

where

$$E(\boldsymbol{\varepsilon}) = \mathbf{0} \quad \text{and} \quad \text{Var}(\boldsymbol{\varepsilon}) = \sigma^2 \mathbf{V}$$

for some $\sigma^2 > 0$ and some known positive definite matrix \mathbf{V} .

Because $\sigma^2 \mathbf{V}$ is a variance matrix, \mathbf{V} is symmetric and positive definite, $\therefore \exists$ a symmetric and positive definite matrix $\mathbf{V}^{1/2} \ni$

$$\mathbf{V}^{1/2} \mathbf{V}^{1/2} = \mathbf{V} \quad \text{and} \quad \mathbf{V}^{1/2} \text{ is nonsingular with } \mathbf{V}^{-1/2} \equiv (\mathbf{V}^{1/2})^{-1}.$$

It follows that under the AM,

$$\mathbf{V}^{-1/2}\mathbf{y} = \mathbf{V}^{-1/2}\mathbf{X}\boldsymbol{\beta} + \mathbf{V}^{-1/2}\boldsymbol{\varepsilon} \iff \mathbf{z} = \mathbf{U}\boldsymbol{\beta} + \boldsymbol{\delta},$$

where

$$\mathbf{z} = \mathbf{V}^{-1/2}\mathbf{y}, \quad \mathbf{U} = \mathbf{V}^{-1/2}\mathbf{X}, \quad \text{and} \quad \boldsymbol{\delta} = \mathbf{V}^{-1/2}\boldsymbol{\varepsilon}$$

with

$$E(\boldsymbol{\delta}) = \mathbf{0}$$

and

$$\begin{aligned}\text{Var}(\boldsymbol{\delta}) &= \mathbf{V}^{-1/2}\sigma^2\mathbf{V}\mathbf{V}^{-1/2} \\ &= \sigma^2\mathbf{V}^{-1/2}\mathbf{V}^{1/2}\mathbf{V}^{1/2}\mathbf{V}^{-1/2} \\ &= \sigma^2\mathbf{I}.\end{aligned}$$

Thus, the AM for \mathbf{y} is equivalent to the GMM for $\mathbf{z} = \mathbf{V}^{-1/2}\mathbf{y}$.

Estimability in the AM:

The AM is just a special case of the GLM.

Thus, as before, $c'\beta$ is estimable iff $c \in \mathcal{C}(X')$.

Note that

$$\begin{aligned}\mathcal{C}(\mathbf{X}') &= \mathcal{C}(\mathbf{X}'\mathbf{V}^{-1/2}) \\ &= \mathcal{C}((\mathbf{V}^{-1/2}\mathbf{X})') \\ &= \mathcal{C}(\mathbf{U}').\end{aligned}$$

Thus, $\mathbf{c} \in \mathcal{C}(\mathbf{X}') \iff \mathbf{c} \in \mathcal{C}(\mathbf{U}')$.

Let \mathcal{L}_y be the collection of all linear estimators that are linear in y .

Let \mathcal{L}_z be the collection of all linear estimators in $z = V^{-1/2}y$. Show that

$$\mathcal{L}_y = \mathcal{L}_z.$$

Estimating $E(\mathbf{y})$ under the Aitken Model:

Consider $Q_{\text{GLS}}(\mathbf{b}) = (\mathbf{y} - \mathbf{X}\mathbf{b})'\mathbf{V}^{-1}(\mathbf{y} - \mathbf{X}\mathbf{b})$.

Finding $\hat{\boldsymbol{\beta}}_{\text{GLS}}$ that minimizes $Q_{\text{GLS}}(\mathbf{b})$ over $\mathbf{b} \in \mathbb{R}^p$ is a Generalized Least Squares (GLS) problem.

If

$$Q_{\text{GLS}}(\hat{\beta}_{\text{GLS}}) \leq Q_{\text{GLS}}(\mathbf{b}) \quad \forall \mathbf{b} \in \mathbb{R}^p,$$

$\hat{\beta}_{\text{GLS}}$ is a solution to the GLS problem.

$X\hat{\beta}_{\text{GLS}}$ is known as GLS estimator of $E(\mathbf{y})$ if $\hat{\beta}_{\text{GLS}}$ is a solution to the GLS problem.

Show that $\hat{\beta}_{\text{GLS}}$ minimizes $Q_{\text{GLS}}(\mathbf{b})$ over $\mathbf{b} \in \mathbb{R}^p$ iff $\hat{\beta}_{\text{GLS}}$ solves

$$\mathbf{X}'\mathbf{V}^{-1}\mathbf{X}\mathbf{b} = \mathbf{X}'\mathbf{V}^{-1}\mathbf{y}.$$

These equations are known as the Aitken Equations (AE).

Henceforth, we will use $\hat{\beta}_{\text{GLS}}$ to denote a solution to the AE.

We will use $\hat{\beta}$ or $\hat{\beta}_{\text{OLS}}$ to denote a solution to the NE

$$\mathbf{X}'\mathbf{X}\mathbf{b} = \mathbf{X}'\mathbf{y}. \quad (\text{Ordinary Least Squares})$$

Because of the equivalence between the AE

$$\mathbf{X}'\mathbf{V}^{-1}\mathbf{X}\mathbf{b} = \mathbf{X}'\mathbf{V}^{-1}\mathbf{y}$$

and the NE

$$\mathbf{U}'\mathbf{U}\mathbf{b} = \mathbf{U}'\mathbf{z},$$

we know a solution to AE is

$$(\mathbf{U}'\mathbf{U})^{-1}\mathbf{U}'\mathbf{z} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y}.$$

Theorem 4.2 (Aitken Theorem):

Suppose the Aitken Model holds. If $c'\beta$ is estimable, then $c'\hat{\beta}_{\text{GLS}}$ is the BLUE of $c'\beta$.

Suppose $c'\beta$ is estimable.

Suppose the AM holds.

Find $\text{Var}(c'\hat{\beta}_{\text{GLS}})$.

Estimation of σ^2 under the Aitken Model:

An unbiased estimator of σ^2 is $\frac{z'(I-P_U)z}{n-r}$ based on our previous result for the GMM.

Now, note that

$$\begin{aligned} \mathbf{z}'(\mathbf{I} - \mathbf{P}_U)\mathbf{z} &= \mathbf{z}'(\mathbf{I} - \mathbf{P}_U)'(\mathbf{I} - \mathbf{P}_U)\mathbf{z} \\ &= \|(\mathbf{I} - \mathbf{P}_U)\mathbf{z}\|^2 = \|\mathbf{z} - \mathbf{P}_U\mathbf{z}\|^2 \\ &= \|\mathbf{z} - \mathbf{U}(\mathbf{U}'\mathbf{U})^{-1}\mathbf{U}'\mathbf{z}\|^2 \\ &= \|\mathbf{z} - \mathbf{U}\hat{\boldsymbol{\beta}}_{\text{GLS}}\|^2 = \|\mathbf{V}^{-1/2}\mathbf{y} - \mathbf{V}^{-1/2}\mathbf{X}\hat{\boldsymbol{\beta}}_{\text{GLS}}\|^2 \\ &= \|\mathbf{V}^{-1/2}(\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}_{\text{GLS}})\|^2 \\ &= (\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}_{\text{GLS}})'\mathbf{V}^{-1}(\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}_{\text{GLS}}). \end{aligned}$$

Thus,

$$\hat{\sigma}_{\text{GLS}}^2 \equiv \frac{(\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}_{\text{GLS}})' \mathbf{V}^{-1} (\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}_{\text{GLS}})}{n - r}$$

is an unbiased estimator of σ^2 under the AM.

A Simple Example

Suppose for $i = 1, \dots, n$,

$$y_i = \beta x_i + \varepsilon_i,$$

where $\varepsilon_1, \dots, \varepsilon_n$ are uncorrelated, $E(\varepsilon_i) = 0$ and $\text{Var}(\varepsilon_i) = \sigma^2 x_i > 0$.

Find the BLUE of β and an unbiased estimator of σ^2 .

Find $\text{Var}(\hat{\beta}_{\text{GLS}})$ for this example.

Find $\hat{\beta}_{OLS}$ for this example.

Find $\text{Var}(\hat{\beta}_{\text{OLS}})$ in this example.